

**Comparison of Large-scale Multiple Testing  
Procedures with Two Examples of  
Microarrays**

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# 1 Introduction

## 1.1 The Problem

- Let  $p$ -dimensional random vectors

$$\mathbf{x}_{i1}, \mathbf{x}_{i2}, \dots, \mathbf{x}_{iN_i} \stackrel{iid}{\sim} N_p(\boldsymbol{\mu}_i, \Sigma), \quad i = 1, 2,$$

- mean vectors  $\boldsymbol{\mu}_i = (\mu_{i1}, \mu_{i2}, \dots, \mu_{ip})'$ ,  $i = 1, 2 : p \times 1$
- covariance matrix  $\Sigma = (\sigma_{ij}) : p \times p$ ,  $\Sigma > 0$  (positive definite)

- Interested in testing the multiple hypotheses that

$$H_i : \mu_{1i} = \mu_{2i} \quad \text{vs.} \quad A_i : \mu_{1i} \neq \mu_{2i}, \quad i = 1, 2, \dots, p,$$

- Assume  $p$  is very large, particularly  $N = N_1 + N_2 < p$  (number of observations smaller than the dimension)

## 1.2 Motivation

- DNA microarrays ( $N < p$ ): thousands of gene expression values are measured on relatively fewer subjects.
  - Example: 6817 ( $p$ ) genes vs. 72 ( $N$ ) subjects in Dudoit, et. al. (2002)
- When  $p$  is large, classical testing procedures are TOO CONSERVATIVE.
  - Example: Bonferroni procedure rejects each  $H_j$  at significance level of  $\alpha/p$  such that

$$\begin{aligned} FWER &= \Pr\{at\ least\ one\ H_j\ is\ falsely\ rejected\} \\ &\leq \sum_{j=1}^p \Pr\{H_j\ is\ falsely\ rejected\} = \alpha \end{aligned}$$

- When  $p$  is large, FWER could be much smaller than  $\alpha$ . Bonferroni procedure is too conservative.

## 2 Large-scale Multiple Testing Procedures.

### 2.1 Notations

- The sample mean vectors

$$\bar{\mathbf{x}}_i = N_i^{-1} \sum_{j=1}^{N_i} \mathbf{x}_{ij} = (\bar{x}_{i1}, \bar{x}_{i2}, \dots, \bar{x}_{ip})', \quad i = 1, 2$$

- $\bar{x}_{ij}$ : the sample mean of the  $j$ th component of the  $i$  group
- The pooled sample covariance matrix

$$S = n^{-1} \sum_{i=1}^2 \sum_{j=1}^{N_i} (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)(\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' = (s_{ij}), \quad n = N - 2$$

- $s_{ij}$ : the pooled sample covariance of the  $i$ th and the  $j$ th component;
- $s_{jj}$ : the common sample variance of the  $j$ th component.

- The test statistic for the single hypothesis  $H_j : \mu_{1j} = \mu_{2j}$

$$T_j = \frac{N_1 N_2}{(N_1 + N_2)} \frac{(\bar{x}_{1j} - \bar{x}_{2j})^2}{s_{jj}} \sim F_{1,n}, \text{ under } H_j$$

- $F_{1,n}$  is the F-distribution with 1 and n degrees of freedom
- $n = N_1 + N_2 - 2 = N - 2$ .

- The p-value of  $T_j$  when  $T_j = t_j$ :

$$p_j = \text{P}\{T_F > t_j | H_j : \mu_{1j} = \mu_{2j}\}$$

- $T_F$  is an  $F_{1,n}$  random variable

- Order the p-values  $p_1, p_2, \dots, p_p$  and the corresponding hypotheses  $H_1, H_2, \dots, H_p$  as

$$p_{(1)} \leq p_{(2)} \leq \dots \leq p_{(p)}$$

$$H_{(1)}, H_{(2)}, \dots, H_{(p)}$$

## 2.2 Type I Error Rates

Let  $V$  be the number of false rejections and  $R$  be the total number of rejections.

- Family-wise Error Rate (FWER)

$$P\{V \geq 1\}.$$

That is, the rate of at least one false rejection of a true null hypothesis.

- False Discovery Rate (FDR)

$$E[V/R] = E[V/R | R > 0]P(R > 0).$$

That is the expected proportion of false rejections among all rejections.

- $\gamma$ FWER

$$P\{V \geq \gamma \times p\},$$

where  $1/p \leq \gamma < 1$ . That is the rate of at least  $100\gamma\%$  false rejections among all the hypotheses.

- Connections between FWER, FDR, and  $\gamma$ FWER

$$FWER \geq FDR$$

$FWER = FDR$ , if all  $H_j$ ,  $j = 1, 2, \dots, p$ , are true.

$$FWER \geq \gamma FWER$$

$$FWER = \gamma FWER, \text{ if } \gamma = 1/p$$

Therefore,

- Control of FWER is MORE STRINGENT than control of FDR or  $\gamma$ FWER
- Control of FDR or  $\gamma$ FWER without controlling FWER may potentially INCREASE the POWER of the testing procedure.



## 2.3 Large-scale Multiple Testing Procedures

- The Procedures

| Method         | Reject   | Control*      | Restriction       |
|----------------|--|---------------|-------------------|
| Bonf           | $H_j, \text{ if } p_j \leq \alpha/p$   | FWER          | No                |
| Holm           | $H_{(j)}, \text{ if } \forall i \leq j, p^{(i)} \leq \frac{\alpha}{(p-i+1)}$             | FWER          | No                |
| Hochberg       | $H_{(j)}, \text{ if } \exists i \geq j, p^{(i)} \leq \frac{\alpha}{(p-i+1)}$             | FWER          | ind. or pos dep** |
| FDRp           | $H_{(j)}, \text{ if } \exists i \geq j, p^{(i)} \leq \frac{j\alpha}{p}$                  | FDR           | ind or pos dep**  |
| cFDRp          | $H_{(j)}, \text{ if } \exists i \geq j, p^{(i)} \leq \frac{j\alpha}{p \sum_{k=1}^j 1/k}$ | FDR           | No                |
| $\gamma$ FWERp | $H_j, \text{ if } p_j \leq \gamma\alpha$ ***   | $\gamma$ FWER | No                |

\* controls FWER or FDR at nominal level  $\alpha$ .

\*\* independence or positive dependence b.w. test statistics  $T_j, j = 1, \dots, p$

\*\*\*  $\frac{1}{p} \leq \gamma < 1$  is preselected, e.g.  $\gamma=0.05; \gamma = 1/p, \gamma$ FWERp  $\iff$  Bonf.

- Conservativeness

$P1 \prec P2$  denotes procedure  $P1$  is more conservative than  $P2$ .

$Bonf \prec Holm \prec Hochberg \prec FDR$

$cFDR \prec FDR$

$Bonf \prec \gamma FWER, \text{ unless } \gamma = 1/p$

$FDR \text{ or } cFDR \text{ ? } \prec \gamma FWER \text{ (depends the value of } \gamma)$

- Less conservative procedure may potentially increase the power at cost of higher Type I error (false rejection rate).

# 3 Examples

## 3.1 The Datasets

- **Colon Data.** 2000 ( $p$ ) gene expression levels are available on 22 ( $N_1$ ) normal colon tissues and 40 ( $N_2$ ) tumor colon tissues. ( $N = 62 < p$ ) [<http://microarray.princeton.edu/oncology/affydata/index.html>; Alon et al. (1999)]

**Q:** Are the tumor genes differentially expressed from the normal genes?

- **Leukemia Data.** 3571 genes expressions are available from 47 ( $N_1$ ) patients suffering from acute lymphoblastic leukemia (ALL) and 25 ( $N_2$ ) patients suffering from acute myeloid leukemia (AML) ( $N = 72 < p$ ) [<http://www.broad.mit.edu/cgi-bin/cancer/datasets.cgi>]; Dudoit, Fridlyland and Speed (2002), Golub et al. (1999)]

**Q:** Are the genes of these two types of cancers differentially expressed?

## 3.2 Results

**Table 3.1** Number of Selected Genes  
Which are Considered as Differentially Expressed

| Colon Dataset ( $p = 2000$ )              |     |      |       | Leukemia Dataset ( $p = 3571$ )           |      |      |       |
|---|-----|------|-------|---|------|------|-------|
| $\alpha$                                  | 0.1 | 0.05 | 0.025 | $\alpha$                                  | 0.1  | 0.05 | 0.025 |
| Bonf <sup>1</sup>                         | 70  | 53   | 45    | Bonf <sup>1</sup>                         | 306  | 279  | 228   |
| Holm <sup>1</sup>                         | 70  | 55   | 45    | Holm <sup>1</sup>                         | 311  | 284  | 233   |
| Hoch <sup>1</sup>                         | 70  | 55   | 45    | Hoch <sup>1</sup>                         | 311  | 284  | 233   |
| FDR <sub>p</sub> <sup>2</sup>             | 478 | 354  | 257   | FDR <sub>p</sub> <sup>2</sup>             | 1366 | 1105 | 912   |
| cFDR <sub>p</sub> <sup>2</sup>            | 188 | 143  | 95    | cFDR <sub>p</sub> <sup>2</sup>            | 736  | 607  | 511   |
| $\gamma_1$ FWER <sub>p</sub> <sup>3</sup> | 292 | 243  | 193   | $\gamma_1$ FWER <sub>p</sub> <sup>3</sup> | 865  | 746  | 644   |
| $\gamma_2$ FWER <sub>p</sub> <sup>3</sup> | 369 | 292  | 243   | $\gamma_2$ FWER <sub>p</sub> <sup>3</sup> | 991  | 865  | 746   |

1:  $FWER \leq \alpha$ ; 2:  $FDR \leq \alpha$ ; 3:  $\gamma$ - $FWER \leq \alpha$ ;  $\gamma_1 = 0.05$ ,  $\gamma_2 = 0.1$ .

**Table 3.2** Number of Selected Genes  
Which are Considered as Differentially Expressed  
with Total Number of Genes Having Been Reduced

| Colon Dataset                             |     |     |     |      | Leukemia Dataset                          |      |      |     |      |
|---|-----|-----|-----|------|---|------|------|-----|------|
| Reduced $p$                               | 478 | 354 | 257 | 2000 | Reduced $p$                               | 1366 | 1105 | 912 | 3571 |
| Bon <sup>1</sup>                          | 85  | 94  | 101 | 53   | Bon <sup>1</sup>                          | 320  | 332  | 341 | 279  |
| Holm <sup>1</sup>                         | 92  | 102 | 139 | 55   | Holm <sup>1</sup>                         | 336  | 356  | 387 | 284  |
| cFDR <sub>p</sub> <sup>2</sup>            | 280 | 326 | 257 | 143  | cFDR <sub>p</sub> <sup>2</sup>            | 818  | 874  | 912 | 607  |
| $\gamma_1$ FWER <sub>p</sub> <sup>3</sup> | 243 | 243 | 243 | 243  | $\gamma_1$ FWER <sub>p</sub> <sup>3</sup> | 745  | 745  | 747 | 746  |
| $\gamma_2$ FWER <sub>p</sub> <sup>3</sup> | 292 | 292 | 257 | 292  | $\gamma_2$ FWER <sub>p</sub> <sup>3</sup> | 865  | 863  | 864 | 865  |

- 1:  $FWER \leq 0.05$ ;
- 2:  $FDR \leq 0.05$ ;
- 3:  $\gamma FWER \leq 0.05$ ;  $\gamma_1 = 0.05$ ,  $\gamma_2 = 0.1$ .

# References

- [1] ALON, U., BARKAI, N., NOTTERMAN, D. A., GISH, K., YBARRA, S., MACK, D. AND LEVINE, A. J. (1999). Broad patterns of gene expression revealed by clustering analysis of tumor and normal colon tissues probed by oligonucleotide arrays. *Proc. Natl. Acad. Sci. U.S.A.* **96** 6745-6750.
- [2] BENJAMINI, Y. AND HOCHBERG, Y. (1995). Controlling the false discovery rate: a practical and powerful approach to multiple testing. *J. Roy. Statist. Soc. B* **57** 289-300.
- [3] BENJAMINI, Y. AND YEKUTIELI, D. (2001). The control of the false discovery rate in multiple testing under dependency. *Ann. Statist.* **4** 1165-1181.
- [4] DUDOLF, S., FRIDLYAND, J. AND SPEED, T. P. (2002). Comparison of discrimination methods for the classification of tremors using gene expression data. *J. Amer. Statist. Assoc.* **97** 77-87.

- [5] DUDOIT, S., SHAFFER J. P. AND BOLDRICK, J. C. (2003). Multiple hypothesis testing in microarray experiments. *Statistical Science* **18** 71-103.
- [6] GOLUB, T. R., SLONIM, D. K., TAMAYO, P., HUARD, C., GAASENBEEK, M., MESIROV, J. P., COLLER, H., LOH, M., DOWNING, J. R., CALIGURI, M. A., BLOOMFIELD, C. D. AND LANDER, E. S. (1999). Molecular classification of cancer: Class discovery and class prediction by gene expression monitoring. *Science* **286** 531-537.
- [7] HOMMEL, G. AND HOFFMAN, T. (1988). Controlled uncertainty. In *Multiple Hypothesis Testing* (P. Bauer, G. Hommel and E. Sonnemann, eds) 154-161. Springer, Heidelberg.
- [8] LEHMANN, E. L. AND ROMANO, J. P. (2005). Generalisations of the familywise error rate. *Ann. Statist.* **33** 1138-1154.
- [9] SIMES, R. J. (1986). An improved Bonferroni procedure for multiple tests of significance. *Biometrika* **73** 751-754.