

Bayesian LAGO for Statistical Detection Problems

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Outline

- The statistical detection problem.
- LAGO—An adaptive detection method.
- BLAGO—Bayesian LAGO.
- Conclusions and discussions.

The Statistical Detection Problem

- Set-up: $\{\mathbf{x}_i, y_i\}_{i=1}^n$, $\mathbf{x}_i \in \mathbb{R}^d$ is a vector of predictors,
 $y_i \in \{0, 1\}$ is class label. (binary classification?).
- Class frequencies: **extremely unbalanced**.
- Objective: Detect those class-1 observations rather than
assign class label to new data.

Performance Measure

- Misclassification rate (MR) is not proper.

True label	classification result	
	0	1
0	a (true negative)	b (false positive)
1	c (false negative)	d (true positive)

And the misclassification rate (MR) is given by

$$\text{MR} = \frac{b + c}{a + b + c + d}$$

- For detection problem, Average Precision (AP) is more suitable.

$$\text{AP} = \frac{1}{n_1} \sum_{i=1}^n \frac{y_{(i)} h(i)}{i}$$

Examples

- Google search. Prefer searching algorithms that return the relevant documents as early as possible.
 $(1,0) = (\text{relevant}, \text{irrelevant})$.

- Drug discovery. Prefer a models that detect those active drug agent for a certain disease as soon as possible.
 $(1,0) = (\text{active compound}, \text{inactive compound})$.

LAGO: An Adaptive Method

- LAGO—Locally Adjusted GO estimate.

- By Bayes' rule:

$$\begin{aligned} p(Y = 1 | \mathbf{z}) &= \frac{p(\mathbf{z}|Y = 1)p(Y = 1)}{p(\mathbf{z}|Y = 1)p(Y = 1) + p(\mathbf{z}|Y = 0)p(Y = 0)} \\ &= \frac{\alpha f(\mathbf{z})}{\alpha f(\mathbf{z}) + 1}, \end{aligned} \tag{1}$$

where $\alpha = \frac{p(Y=1)}{p(Y=0)}$ and $f(\mathbf{z}) = \frac{p(\mathbf{z}|Y=1)}{p(\mathbf{z}|Y=0)}$.

Step 1 in Estimating $f(\mathbf{z})$

- Estimate $p(\mathbf{z}|Y = 1)$ by kernel density estimate

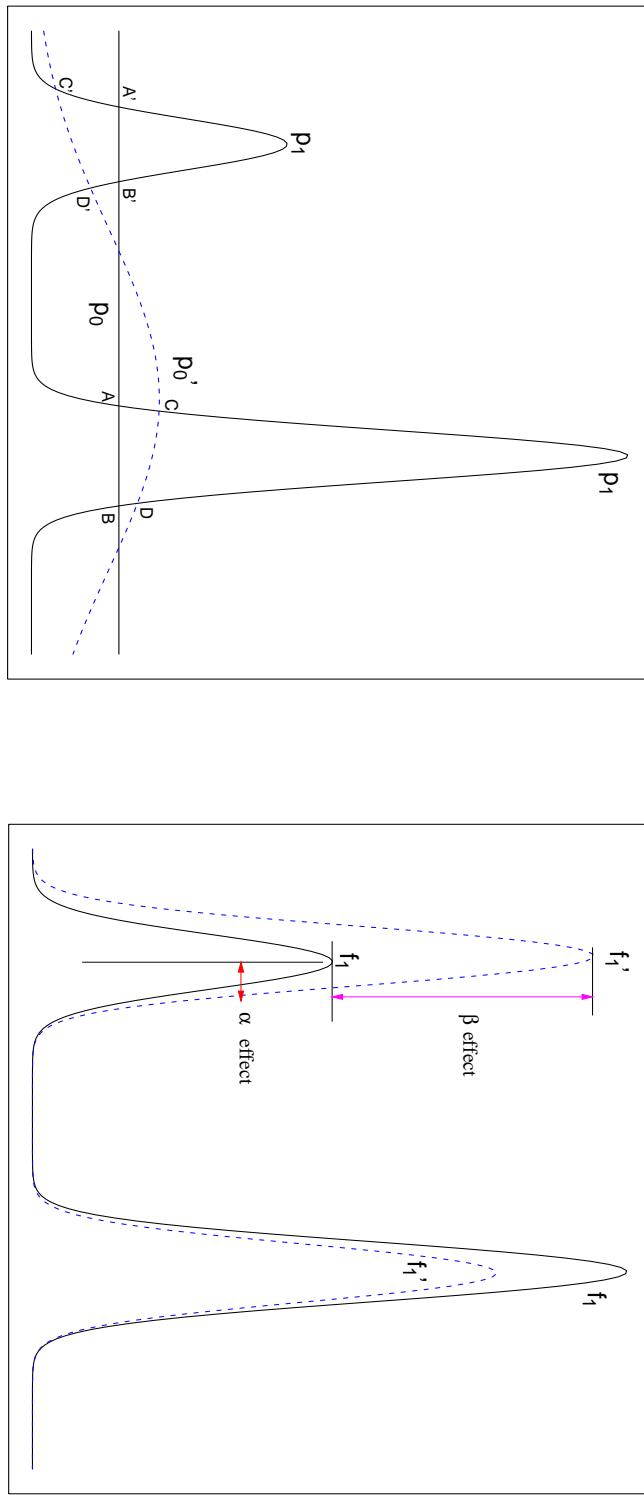
$$p(\mathbf{z}|Y = 1) = \frac{1}{N_1} \sum_{i=1}^{N_1} \prod_{j=1}^d \frac{1}{r_{ij}} g\left(\frac{z_j - x_{ij}}{r_{ij}}\right),$$

where $r_{ij} = \frac{1}{K} \sum_{\mathbf{w} \in N(\mathbf{x}_i, K)} |w_{ij} - x_{ij}|$ is the average distance in the j -th dimension between a class-1 observation \mathbf{x}_i and its K nearest **class-0** neighbors

- If $p(\mathbf{z}|Y = 0) = C$, $p(\mathbf{z}|Y = 1) \propto f(\mathbf{z})$, and hence $p(\mathbf{z}|Y = 1) \propto p(Y = 1|\mathbf{z})$.

Examining the ripple effects

Note: Left: The density functions of p_0 and p_1 . Right: The posterior distributions.



Step 2 in Estimating $f(\mathbf{z})$

- Consider the ripple effects, adjust $p(\mathbf{z}|Y = 1)$ according to the local density of the class-0 observations, we obtain the LAGO estimate

$$\begin{aligned}\hat{f}(\mathbf{z}) &= \frac{1}{N_1} \sum_{i=1}^{N_1} \prod_{j=1}^d \frac{1}{\alpha r_{ij}} r_{ij}^\beta g\left(\frac{z_j - x_{ij}}{\alpha r_{ij}}\right) \\ &\propto \frac{1}{N_1} \sum_{i=1}^{N_1} \prod_{j=1}^d g\left(\frac{z_j - x_{ij}}{\alpha r_{ij}}\right).\end{aligned}\quad (2)$$

- LAGO paper (2006): Technometrics, 48(2), 193-205.

Motivation of BLAGO

- Only provides a point estimate, no uncertainty is embodied.
- LAGO score is within $(0,1)$, but not a probability.
- A more precise estimate.

Bayesian Methods in Machine Learning

- Bayesian Neural Network (Dr. Radford Neal)
- Bayesian Tree (Dr. Hugh Chipman)
- Bayesian K-nearest-neighbor (Dr. Chris Holmes)

Likelihood Function

- LAGO score is not probabilistic, change it to probability scale by a logistic transformation

$$\theta_i = p(y_i = 1 | \mathbf{z}_i) = \frac{e^{\beta_0 + \beta_1 \hat{f}(\mathbf{z}_i)}}{1 + e^{\beta_0 + \beta_1 \hat{f}(\mathbf{z}_i)}}, \quad (3)$$

where (\mathbf{z}_i, y_i) is the i th test point and $\beta_1 > 0$.

- The pseudo-likelihood is

$$L(\boldsymbol{\theta}; \mathbf{Y}) = \prod_{i=1}^n \theta_i^{y_i} [1 - \theta_i]^{1-y_i}. \quad (4)$$

Prior Distributions

- $K \sim \text{UNIF}(1, 2, \dots, K_{\max})$.
- $\alpha \sim \text{UNIF}(0, \infty)$.
- $(\beta_0, \beta_1) \sim \text{BVN}(\mu_0, \Sigma_0)$ with $\beta_1 > 0$.

MCMC Steps

Step 0. Set the initial K_{old} , α_{old} ;

Step 1. Propose new K_{new} and α_{new} by

$$K_{new} = K_{old} \pm U[0, 1, 2, 3], \quad \alpha_{new} = \alpha_{old} + N(0, \sigma_\alpha^2).$$

Step 2. Generate a random number $\gamma \sim U(0, 1)$. Accept the new state K_{new} , α_{new} if $\gamma \leq J$ with accepted probability

$$J = \min \left\{ \frac{p(K_{new}, \alpha_{new} | \mathbf{Y})}{p(K_{old}, \alpha_{old} | \mathbf{Y})}, 1 \right\}. \quad (5)$$

Step 3. Repeat steps 2 and 3 until burn-in.

Laplace Approximation

The well-known Laplace approximation is given by

$$\begin{aligned} I_n &= \int_a^b b(\hat{\beta}) \exp\{-nr(\hat{\beta})\} dt \\ &\approx b(\hat{\beta}) \left(\frac{2\pi}{n}\right)^{\frac{p}{2}} (\det \Sigma_{\hat{\beta}})^{\frac{1}{2}} e^{-nr(\hat{\beta})}, \end{aligned} \quad (6)$$

where p is the dimension of β ; and $\hat{\beta}$ is the solution to

$$\nabla r(\hat{\beta}) = 0; \text{ and } \Sigma_{\hat{\beta}} = [r''(\hat{\beta})]^{-1}.$$

Evaluating $p(K, \alpha | \mathbf{Y})$

$$\begin{aligned} p(K, \alpha | \mathbf{Y}) &\propto p(\mathbf{Y} | K, \alpha) \pi(K, \alpha) \\ &\propto \int p(\mathbf{Y} | K, \alpha, \boldsymbol{\beta}) \pi(\boldsymbol{\beta} | K, \alpha) d\boldsymbol{\beta} \\ &= \int L(\boldsymbol{\beta}) \pi(\boldsymbol{\beta}) d\boldsymbol{\beta} \\ &= \int \exp(-n(-\frac{1}{n} \mathcal{L}(\boldsymbol{\beta}))) d\boldsymbol{\beta} \\ &= \int \exp\{-n(-\frac{1}{n} \mathcal{L}(\boldsymbol{\beta}))\} d\boldsymbol{\beta} \\ &\approx (2\pi)^{p/2} e^{\mathcal{L}(\boldsymbol{\beta}^* | K, \alpha)} |H(\boldsymbol{\beta}^* | K, \alpha)|^{-1/2}. \end{aligned} \quad (7)$$

- $\mathcal{L}(\boldsymbol{\beta}^*)$: Log posterior of $\boldsymbol{\beta}$ evaluated at the posterior mode $\boldsymbol{\beta}^*$
- $H(\boldsymbol{\beta}^*)$: Hessian matrix of the log posterior of $\boldsymbol{\beta}$ evaluated at $\boldsymbol{\beta}^*$.

Prediction

Suppose

$$\theta(\mathbf{z}_i|\text{train}, k_0, \alpha_0, \boldsymbol{\beta}_0) = \frac{e^{\mathbf{x}_i^T \boldsymbol{\beta}_0}}{1 + e^{\mathbf{x}_i^T \boldsymbol{\beta}_0}}, \quad (8)$$

where $\mathbf{x}_i = [1, \hat{f}(\mathbf{z}_i|\text{train}, k_0, \alpha_0)]^T$. Incorporating uncertainty, the prediction

$$\begin{aligned} \theta(\mathbf{z}_i|\text{train}) &\approx \frac{1}{M} \sum_{j=1}^M \theta(\mathbf{z}_i|\text{train}, k^{(j)}, \alpha^{(j)}) \\ &= \frac{1}{M} \sum_{j=1}^M \int \theta(\mathbf{z}_i|\text{train}, k^{(j)}, \alpha^{(j)}, \boldsymbol{\beta}) p(\boldsymbol{\beta}|\text{train}, k^{(j)}, \alpha^{(j)}) d\boldsymbol{\beta}, \end{aligned}$$

where $k^{(j)}$ and $\alpha^{(j)}$ are the j th ($j = 1, \dots, M$) sample in the converged chain.

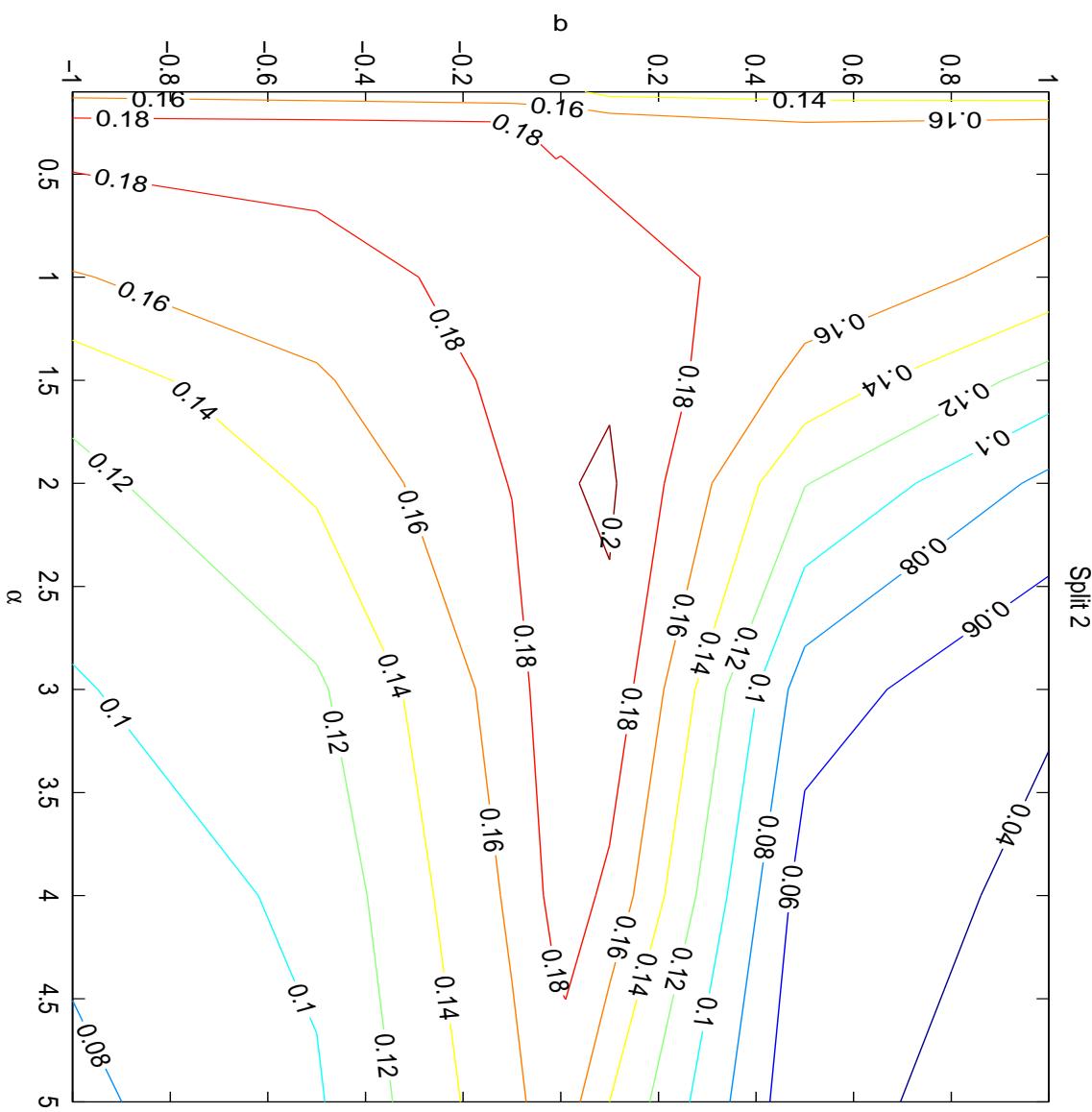
Conclusions and Discussions

- Strength: BLAGO provides a probabilistic prediction and uncertainty of the prediction.
- Challenges:
 - Choice of the prior distributions.
 - The proposal functions.
 - Performance measure.

Reparameterization

- Consider the ripple effect, α , the kernel function ($d=1$) is
$$\frac{1}{\alpha r_i} g\left(\frac{z - x_i}{\alpha r_i}\right)$$
- Incorporate the other ripple effect, β , the score function is
$$\hat{p}(\mathbf{z}) = \frac{1}{n_1} \sum_{i=1}^{n_1} \left(\frac{1}{\alpha r_i} \right) (\alpha r_i)^\beta g\left(\frac{z - x_i}{\alpha r_i}\right) \propto \frac{1}{n_1} \sum_{i=1}^{n_1} r_i^{\beta-1} g\left(\frac{z - x_i}{\alpha r_i}\right)$$
- An empirical study shows that $b = \beta - 1 \approx 0$.

Average precision maximized around $b \approx 0$



The NCI database

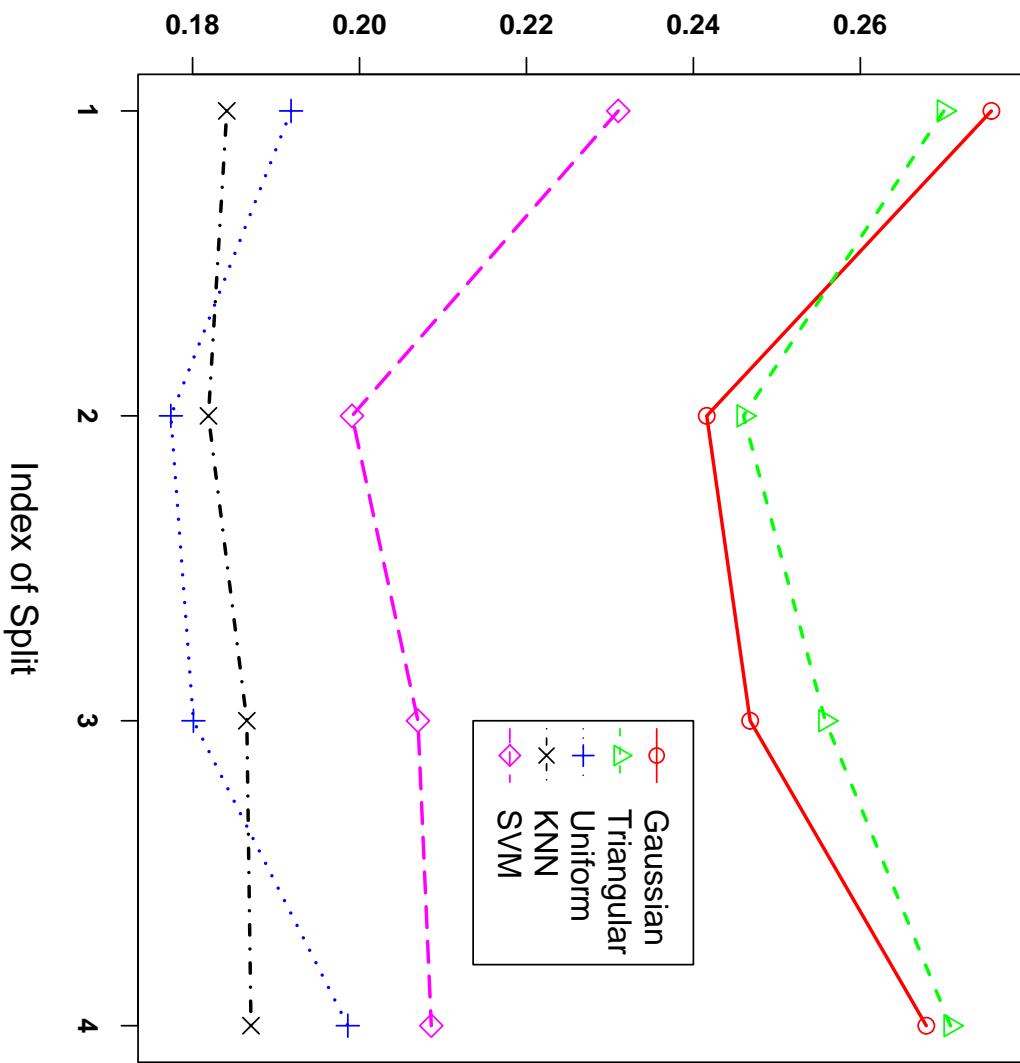
AIDS Antiviral Data—from National Cancer Institute (NCI)

- Target—HIV1 virus;
- Response: {0, 1} or {inactive, active}
- Explanatory variables: Molecular descriptors
(BCUT₁, . . . , BCUT₆).

	Training	Test	
Active	304	304	608
Inactive	14,602	14,602	29,204
Total	14,906	14,906	29,812

Results and Comparisons

Average Precision



A Numeric Summary of Hit Curve

$$\text{Average Precision} = \frac{1}{n_1} \sum_{i=1}^n \frac{y_{(i)} h(i)}{i}$$

Ranking	True Class Label	Precision
1	1	1/1
2	1	2/2
3	0	
4	1	3/4
5	0	
6	0	

$$\text{AP} = \frac{1}{3}(1/1 + 2/2 + 3/4) \approx 0.9167$$